Limit theorems for subcritical age-dependent branching processes with two types of immigration

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 2 Department of Operational Research, Probability and Statistics, University of Sofia

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The effect of the following two-type immigration pattern is studied.

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At a sequence of renewal epochs a random number of immigrants enters the population.

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At a sequence of renewal epochs a random number of immigrants enters the population.

Each subpopulation stemming from one of these immigrants or one of the ancestors is revived by new immigrants and their offspring whenever it dies out, possibly after an additional delay period.

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At a sequence of renewal epochs a random number of immigrants enters the population.

Each subpopulation stemming from one of these immigrants or one of the ancestors is revived by new immigrants and their offspring whenever it dies out, possibly after an additional delay period.

All individuals have the same lifetime distribution and produce offspring according to the same reproduction law. This is the Bellman-Harris process with immigration at zero and immigration of renewal type (BHPIOR). Introduction History and related . . Model description Sample Path of BHIO Sample Path of BHIOR Notations Results References Home Page Title Page **44** •• Page 2 of 22 Go Back Full Screen Close Quit

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We prove a strong law of large numbers and a central limit theorem for such processes.

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Similar conclusions are obtained for their discrete-time counterparts (lifetime per individual equals one), called Galton- Watson processes with immigration at zero and immigration of renewal type (GWPIOR).

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Similar conclusions are obtained for their discrete-time counterparts (lifetime per individual equals one), called Galton- Watson processes with immigration at zero and immigration of renewal type (GWPIOR).

Our approach is based on the theory of regenerative processes, renewal theory and occupation measures and is quite different from those in earlier related work using analytic tools.

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 $\{Z(t)\}_{t\geq 0}$ be a Bellman-Harris branching process with immigration only in the state zero (BHBPIO) with model parameters an individual lifetime distribution G with G(0) = 0, an offspring distribution $(p_j)_{j\geq 0}$ with p.g.f. f(s), a number of immigrants distribution $(g_j)_{j\geq 0}$ with p.g.f. g(s), a distribution D of the delay times elapsing after extinction epochs before new immigrants enter the population.



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The discrete-time variant $(Z(n))_{n\geq 0}$, where $t \in [0,\infty]$ is replaced with $n \in \mathbb{N}_0$, and where $G = \delta_1$ (Dirac measure at 1) and D is a distribution on \mathbb{N}_0 , will be called a Galton-Watson process with immigration at 0 (GWPIO).



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 $Z_{ij} = (Z_{ij}(t))_{t \ge 0}$, $i \ge 0$, $j \ge 1$ - i.i.d. BHPIO with the same model parameters as $\{Z(t)\}_{t \ge 0}$

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 $(\sigma_n)_{n\geq 0}$ - zero-delayed renewal process with increment distribution F

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The numbers of immigrants $(Y_n)_{n\geq 1}$ are assumed to be iid r.v.'s with probability generating function (pgf) h(s). The Y_n are supposed to be the numbers of individuals entering the population at times σ_n . A further integer-valued random variable Y_0 gives the number of ancestors of the considered population. It is assumed that σ_n , $(Y_n)_{n\geq 1}$, Y_0 and all Z_{ij} are mutually independent.



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$$N(t) \stackrel{def}{=} \sup\{n : \sigma_n \le t\}$$

the number of renewal events in the sequence $\sigma_n, n = 1, 2, ...$ during the time interval [0, t].

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A Bellman-Harris process with immigration at zero and immigration of renewal type (BHPIOR) X(t) can be defined as follows

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A Bellman-Harris process with immigration at zero and immigration of renewal type (BHPIOR) X(t) can be defined as follows

$$X(t) \stackrel{def}{=} \sum_{i=0}^{N(t)} Z_i(t - \sigma_i), \quad t \ge 0,$$

and

$$Z_i(t) = \sum_{i=0}^{Y_i} Z_{ij}(t), \quad t \ge 0$$

is a BHPIO with Y_i ancestors.



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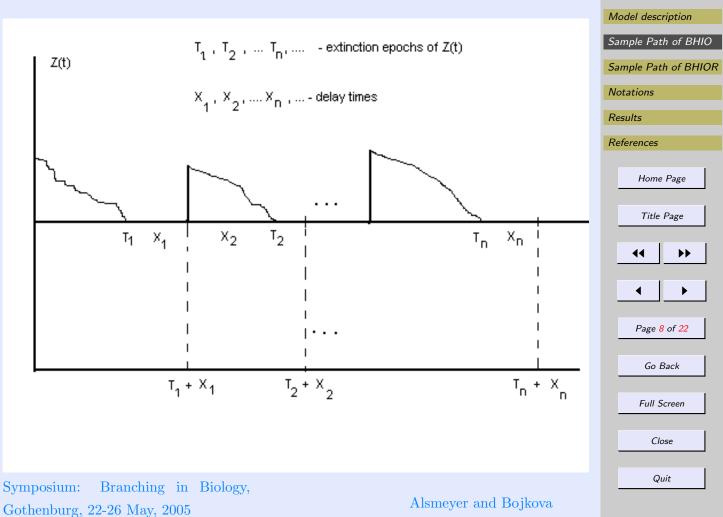
is a BHPIO with Y_i ancestors.

Its discrete time variant, where the Z_i are GWPIO and σ_n forms a discrete renewal process, is called a Galton-Watson process with immigration at zero and immigration of renewal type (GWPIOR).

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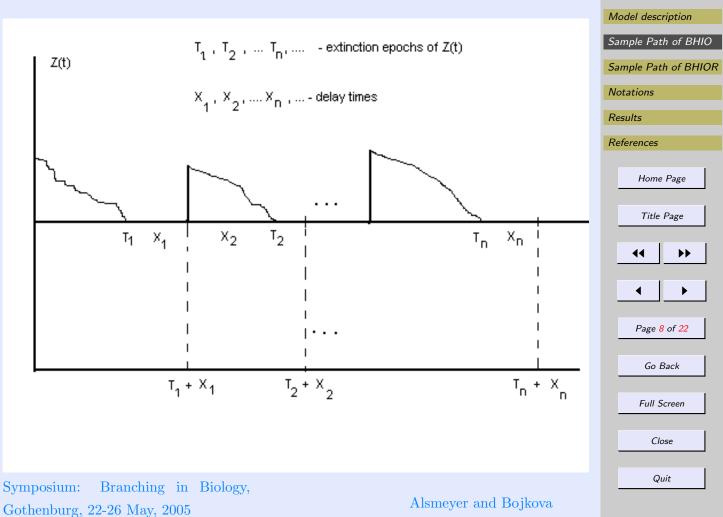
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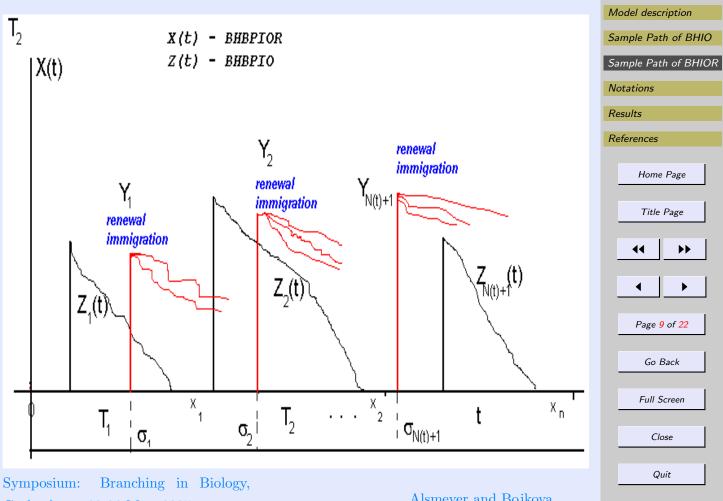
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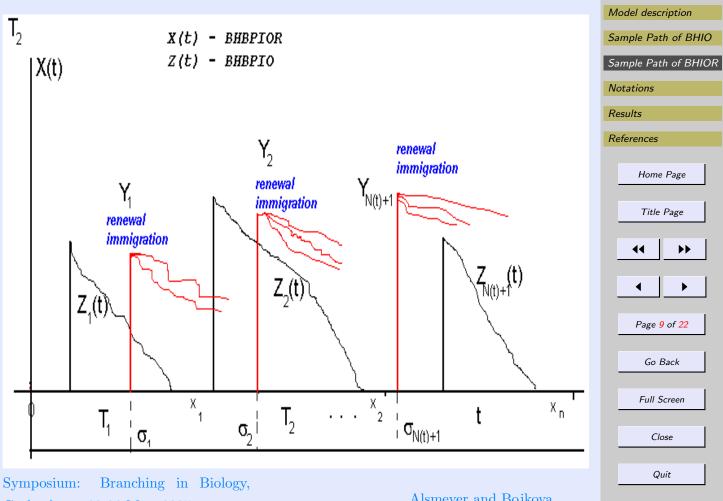
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$$m \stackrel{def}{=} \sum_{k \ge 1} k p_k = f'(1), \quad m_G \stackrel{def}{=} \int_0^\infty t dG(t)$$

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$$m \stackrel{def}{=} \sum_{k \ge 1} k p_k = f'(1), \quad m_G \stackrel{def}{=} \int_0^\infty t dG(t)$$

similarly m_F and m_D .

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$$m \stackrel{def}{=} \sum_{k \ge 1} k p_k = f'(1), \quad m_G \stackrel{def}{=} \int_0^\infty t dG(t)$$

similarly m_F and m_D .

Let the *p*th moments of $(p_k)_{k\geq 0}$, *G*, *F*, *D* be denoted as m_p , $m_{G,p}$, $m_{F,p}$ and $m_{D,p}$, respectively.

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Put $\mathbb{P}_k = \mathbb{P}(.|Z(0) = k)$ for $k \ge 0$ and $\mathbb{P}^* \stackrel{def}{=} \sum_{k\ge 0} g_k \mathbb{P}_k$, so that the initial distribution of $(Z(t))_{t\ge 0}$ under \mathbb{P}^* is $(g_k)_{k\ge 0}$.

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We will simply write \mathbb{P} in assertions where the distribution of Z(0) does not matter.

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Let T_1 be the first extinction epoch of $(Z(t))_{t\geq 0}$ after 0, defined as

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Let T_1 be the first extinction epoch of $(Z(t))_{t\geq 0}$ after 0, defined as

$$T_1 \stackrel{def}{=} \inf\{t > 0 : Z(t-) > 0, Z(t) = 0\}$$

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$T_1 \stackrel{def}{=} \inf\{n \ge 1 : Z(n) = 0\}$

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$$(\widehat{Z}(t))_{t\geq 0} \stackrel{def}{=} (Z(t)\mathbb{I}_{T_1>t})_{t\geq 0}$$

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$$\Phi(s,t) \stackrel{def}{=} \mathbb{E}_1 s^{\widehat{Z}(t)}$$

- the p.g.f. of $\widehat{Z}(t)$ under \mathbb{P}_1 and

$$m(t) \stackrel{def}{=} \mathbb{E}_1 \widehat{Z}(t)$$

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$$\Lambda(t) \stackrel{def}{=} \mathbb{E}^* Z(t) \text{ and } \Lambda_2(t) \stackrel{def}{=} \mathbb{E}^* Z(t)^2.$$

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When moving to the process $(X(t))_{t\geq 0}$ we put $Z(t) \stackrel{def}{=} Z_0(t)$ for $t \geq 0$ and retain the previous notation.

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Proposition 7.1.

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Proposition 7.1.

Let $(Z(t))_{t\geq 0}$ be a subcritical BHPIO with arbitrary ancestor distribution, $g'(1) < \infty$, and $m_G < \infty$. Suppose also $m_D < \infty$, and that the convolution G * D is nonarithmetic.

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Then $Z(t) \xrightarrow{d} Z(\infty)$, $t \to \infty$, for an integer-valued random variable $Z(\infty)$ satisfying

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where $\beta \stackrel{def}{=} \mathbb{E}^*T_1 + m_D$ is finite. $Z(\infty)$ has p.g.f.

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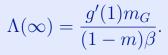
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$$\Phi(s,\infty) = \frac{m_D}{\beta} + \frac{1}{\beta} \int_0^\infty (g(\Phi(s,t)) - g(\Phi(0,t))) dt$$

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$$\Lambda(\infty) = \frac{g'(1)m_G}{(1-m)\beta}$$

Moreover,

$$\lim_{t \to \infty} \mathbb{E}_k Z(t) = \lim_{t \to \infty} \Lambda(t) = \Lambda(\infty)$$



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If f''(1), $m_{G,2}$ and $m_{D,2}$ are all finite, then also

$$\lim_{t \to \infty} \mathbb{E}_k Z^2(t) = \lim_{t \to \infty} \Lambda_2(t) = \Lambda_2(\infty) \stackrel{def}{=} \mathbb{E} Z^2(\infty)$$

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$$\Lambda_2(\infty) = \frac{g'(1)m_G}{(1-m)\beta} + \frac{1}{\beta} \left(\frac{g'(1)f''(1)}{1-m} + g''(1)\right) \int_0^\infty m^2(t)dt < \infty$$

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Let $(X(t))_{t\geq 0}$ be a subcritical BHPIOR with arbitrary ancestor distribution, $g'(1) < \infty$, $h'(1) < \infty$ and $m_G < \infty$.

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Then

$$\frac{X(t)}{t} \xrightarrow{\mathbb{P}} \frac{g'(1)h'(1)m_G}{(1-m)m_F\beta}, \quad t \to \infty$$

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Suppose also $m_F < \infty$, $m_{D,2} < \infty$, and that at least one of G or D is spread out.

Then

$$\frac{X(t) - (N(t) + 1)h'(1)\Lambda(\infty)}{t^{1/2}} \stackrel{d}{\to} N(0, m_F \Xi(\infty)^2)$$

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$$\frac{X(t) - (N(t) + 1)h'(1)\Lambda(\infty)}{t^{1/2}} \xrightarrow{d} N(0, m_F \Xi(\infty)^2)$$

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 $\Xi(\infty)^2 \stackrel{def}{=} (h''(1) - h'(1)^2)\Lambda(\infty)^2 + h'(1)(\Lambda_2(\infty) - \Lambda(\infty)^2)$

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denotes the variance of $Z^*(\infty)$, the limiting variable of $Z_i(t)$.

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Let $(Z(k))_{k\geq 0}$ be a subcritical GWPIO with arbitrary ancestor distribution, $g'(1) < \infty$. Suppose also $m_D < \infty$, and that the convolution $G * D = \delta_1 * D$ is 1-arithmetic.



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$$\Phi(s,\infty) = \frac{m_D}{\beta} + \frac{1}{\beta} \sum_{k \ge 0} (g(\Phi(s,k)) - g(\Phi(0,k)))$$

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$$\Phi(s,\infty) = \frac{m_D}{\beta} + \frac{1}{\beta} \sum_{k \ge 0} (g(\Phi(s,k)) - g(\Phi(0,k)))$$

and mean

$$\Lambda(\infty) = \frac{g'(1)}{(1-m)\beta}$$

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$$\lim_{k \to \infty} \mathbb{E}_j Z(k) = \lim_{k \to \infty} \Lambda(k) = \Lambda(\infty), \quad j \ge 0$$

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$$\frac{X(t)}{t} \xrightarrow{\mathbb{P}} \frac{g'(1)h'(1)}{(1-m)m_F\beta}, \quad t \to \infty$$

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$$\frac{X(t)}{t} \xrightarrow{\mathbb{P}} \frac{g'(1)h'(1)}{(1-m)m_F\beta}, \quad t \to \infty$$

as $t \to \infty$ through the integers.

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 $\frac{X(t) - (N(t) + 1)h'(1)\Lambda(\infty)}{t^{1/2}} \xrightarrow{d} N(0, m_F \Xi(\infty)^2)$



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